

Math 554

Hemanshu Kaul

kaul@iit.edu

The first known application of probabilistic method

Szele 1943  $\exists$  tournament on  $n$  vertices with at least  $n! 2^{-(n-1)}$  Hamiltonian paths.

→ Directed  $K_n$



↳ path of length  $n$

Proof "Create a random tournament" Exercise.  $\square$

• Szele conjectured max # Hamiltonian paths in a tournament on  $n$  vertices is  $n! / (2^{o(n)})$

• Proved by Alon (1990) using Minc-Bregman theorem on permanents (which we will see later while studying Entropy)

Turán's Theorem answers the question

What is the maximum number of edges in an  $n$ -vertex  $K_k$ -free graph?

In the complementary form cliques change to independent sets

So, we are asking this question about graphs without large independent sets.

Coxo (1979) & Wei (1981) showed that a graph with small degrees must contain large independent sets.

This result can be used to give a proof of Turán's Theorem.

Thm (Caro, Wei)

For every graph  $G$ ,

$$\alpha(G) \geq \sum_{v \in V(G)} \frac{1}{d(v)+1}$$

independence #

clique #

Cor  $\omega(G) \geq \sum_{v \in V(G)} \frac{1}{?}$

$d(v) = \text{degree of } v$

Thm (Caro, Wei) For every graph  $G$ ,  $\alpha(G) \geq \sum_{v \in V(G)} \frac{1}{d(v)+1}$

independence #  $\swarrow$

Cor  $\omega(G) \geq \sum_{v \in V(G)} \frac{1}{n-d(v)}$

clique #  $\searrow$

$d(v) = \text{degree of } v$

Cor [Turán's Thm, weaker form]  
Every  $n$ -vertex  $K_{k+1}$ -free graph has at most  $(1 - \frac{1}{k}) \frac{n^2}{2}$  edges

Pf Fill in the details

$$k \geq \omega(G) \geq \sum_{v \in V(G)} \frac{1}{n-d(v)}$$

$\geq ?$  Think!

$$= \frac{n}{(n - \frac{2m}{k})}, \text{ where } m = \# \text{edges}$$

Rearrange to get the needed inequality  $m \leq (1 - \frac{1}{k}) \frac{n^2}{2}$   $\square$

Thm (Caro, Wei) For every graph  $G$ ,  $\alpha(G) \geq \sum_{v \in V(G)} \frac{1}{d(v)+1}$

Proof

- \* Considers a random permutation of the vertices.
- \* Let  $I =$  set of vertices that appear before all its neighbors.

Thm (Caro, Wei) For every graph  $G$ ,  $\alpha(G) \geq \sum_{v \in V(G)} \frac{1}{d(v)+1}$

Proof

- \* Considers a random permutation of the vertices.
- \* Let  $I =$  set of vertices that appear before all its neighbors.

$I$  is an independent set.

For each  $v$ ,  $P[v \in I] = \frac{1}{1+d(v)}$  ( $v$  appears first among  $\{v\} \cup N(v)$ )

$$\therefore E[|I|] = E\left[\sum_{v \in V(G)} I_v\right] = \sum_{v \in V(G)} E[I_v] = \sum_{v \in V(G)} P[v \in I] = \sum_{v \in V(G)} \frac{1}{1+d(v)}$$

$$I_v = \begin{cases} 1 & \text{if } v \in I \\ 0 & \text{if } v \notin I \end{cases}$$

▣

Ques ① Is this bound sharp?

② [DeRandomization] Can you give a deterministic algorithm that gives an independent set of this size?

In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?



In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?

Thm (Alon 1990) Let  $G$  be  $n$ -vertex with  $\delta(G) = d > 1$ .  
Then  $G$  has a dominating set with at most  $\left(\frac{\log(d+1)+1}{d+1}\right)n$  vertices.

Greedy algo?

In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?

Thm (Alon 1990) Let  $G$  be  $n$ -vertex with  $\delta(G) = d > 1$ .  
Then  $G$  has a dominating set with at most  $\left(\frac{\log(d+1)+1}{d+1}\right)n$  vertices.

Proof ① choose a random set ② Add vertices to it to ensure domination

Let  $p \in [0, 1]$  (to be decided later)

Let  $S \subseteq V(G)$  be chosen randomly: each vertex chosen ind. w. prob.  $p$ .

In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?

Thm (Alon 1990) Let  $G$  be  $n$ -vertex with  $\delta(G) = d > 1$ .

Then  $G$  has a dominating set with at most  $\left(\frac{\log(d+1)+1}{d+1}\right)n$  vertices.

Proof ① choose a random set ② Add vertices to it to ensure domination

Let  $p \in [0, 1]$  (to be decided later)

Let  $S \subseteq V(G)$  be chosen randomly: each vertex chosen ind. w. prob.  $p$ .

Let  $T \subseteq V(G) \setminus S$  be the set of all vertices with no neighbors in  $S$ .

Then  $S \cup T$  is a dominating set.

In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?

Thm (Alon 1990) Let  $G$  be  $n$ -vertex with  $\delta(G) = d > 1$ .  
Then  $G$  has a dominating set with at most  $\left(\frac{\log(d+1)+1}{d+1}\right)n$  vertices.

Proof Let  $p \in [0,1]$ .

Let  $S =$  random subset of  $V(G)$  with each vertex chosen ind. with prob.  $p$ .

Let  $T = V(G) \setminus (S \cup N(S))$ .  $\therefore S \cup T$  is dominating.

$$\begin{aligned} \mathbb{E}[|S \cup T|] &= \mathbb{E}[|S| + |T|] = \mathbb{E}[|S|] + \mathbb{E}[|T|] \\ &= \sum_v \mathbb{P}[v \in S] + \sum_v \mathbb{P}[v \in T] \\ &\leq np + n(1-p)^{d+1} \quad \leftarrow \text{Why?} \end{aligned}$$

In a graph  $G=(V,E)$ ,  $U \subseteq V$  is dominating if every vertex in  $V \setminus U$  has a neighbor in  $U$ .

What is the minimum size of a dominating set?

Thm (Alon 1990) Let  $G$  be  $n$ -vertex with  $\delta(G) = d > 1$ .  
Then  $G$  has a dominating set with at most  $\left(\frac{\log(d+1)+1}{d+1}\right)n$  vertices.

Proof Let  $p \in [0,1]$ .

Let  $S =$  random subset of  $V(G)$  with each vertex chosen ind. with prob.  $p$ .

Let  $T = V(G) \setminus (S \cup N(S))$ .  $\therefore S \cup T$  is dominating.

$$E[|S \cup T|] = E[|S| + |T|] = E[|S|] + E[|T|]$$

$$= \sum_v P[v \in S] + \sum_v P[v \in T]$$

$$\leq np + n(1-p)^{d+1} \quad \leftarrow \text{Why?}$$

$$\leq np + ne^{-p(d+1)} \quad \leftarrow \text{Why?}$$

$$\leq n \frac{1 + \log(1+d)}{1+d} \quad \text{for } p = \frac{\log(d+1)}{d+1}$$

Markov Inequality Let  $X \geq 0$  be a random variable.

Then for every  $t > 0$ ,  $P[X \geq t] \leq \frac{E[X]}{t}$ .

$$P_f. E[X] = \sum_{k \geq 0} k P[X=k] \geq \sum_{k \geq t} k P[X=k] \geq t \sum_{k \geq t} P[X=k] = t P[X \geq t].$$

For us,  $X =$  counting variable (discrete r.v.)

so "  $E[X] \rightarrow 0 \Rightarrow P[X=0] \rightarrow 1$  "

Markov Inequality Let  $X \geq 0$  be a random variable.

Then for every  $t > 0$ ,  $P[X \geq t] \leq \frac{E[X]}{t}$ .

How is  $\chi(G)$  determined?

Is it a function of  $\omega(G)$ ? No!

A graph can be very sparse & yet have high  $\chi(G)$ .

Markov Inequality Let  $X \geq 0$  be a random variable.

Then for every  $t > 0$ ,  $P[X \geq t] \leq \frac{E[X]}{t}$ .

How is  $\chi(G)$  determined?

Is it a function of  $\omega(G)$ ? No!

A graph can be very sparse & yet have high  $\chi(G)$ .

Thm [Erdős 1959] Given  $k \geq 3, q \geq 3$ ,  $\exists$  graph with  
girth at least  $q$  and chromatic number at least  $k$ .

$\underbrace{\hspace{10em}}$   
length of shortest cycle



Markov Inequality Let  $X \geq 0$  be a random variable.

Then for every  $t > 0$ ,  $P[X \geq t] \leq \frac{E[X]}{t}$ .

How is  $\chi(G)$  determined?

Is it a function of  $\omega(G)$ ? No!

A graph can be very sparse & yet have high  $\chi(G)$ .

Thm [Erdős 1959] Given  $k \geq 3, q \geq 3$ ,  $\exists$  graph with  
girth at least  $q$  and chromatic number at least  $k$ .

We want a graph with no short cycles  $\rightarrow$  girth  $\geq q$   
and no large independent sets  $\rightarrow \chi(G) \geq \frac{n}{\alpha(G)} \geq k$

We will generate an Erdős-Rényi random graph  $G(n, p)$   
on  $n$  vertices & each pair of vertices is an edge independently  
with probability  $p$ .

Thm [Erdős 1959] Given  $k \geq 3, q \geq 3$ .  $\exists$  graph  $G$  with  
 $girth(G) \geq q$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p) \leftarrow$  with  $n$  to be  
chosen later  
Let  $p = n^{t-1}$  where  $t < 1/q$ , say  $t = 1/2q \leftarrow$  proof works for  
 $\log n/n \ll p \ll n^{1/q-1}$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  
 $girth(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

(# possible cycles of length  $j$  is ?  
and all needed edges are present with probability ?)

Let  $X = \#$  cycles of length  $< g$

$$E[X] = \sum_{j=3}^{g-1} \frac{1}{j} ?$$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $girth(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

(# possible cycles of length  $j$  is  $\binom{n}{j} \frac{(j-1)!}{2}$  and all needed edges are present with probability  $p^j$ )

Let  $X = \#$  cycles of length  $< g$

$$\mathbb{E}[X] = \sum_{j=3}^{g-1} \binom{n}{j} \frac{(j-1)!}{2} p^j \leq \sum n^j p^j = \sum n^j n^{tj-j} = \sum_{j=3}^{g-1} n^{tj} < g n^{1-1/g}$$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $girth(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

(# possible cycles of length  $j$  is  $\binom{n}{j} \frac{(j-1)!}{2}$  and all needed edges are present with probability  $p^j$ )

Let  $X = \#$  cycles of length  $< g$

$$\mathbb{E}[X] = \sum_{j=3}^{g-1} \binom{n}{j} \frac{(j-1)!}{2} p^j \leq \sum n^j p^j = \sum n^j n^{tj-j} = \sum_{j=3}^{g-1} n^{tj} < g n^{1-1/g} = g n^{1/g(g-1)}$$

Since  $g$  is fixed,  $\mathbb{E}[X] \rightarrow 0$  as  $n \rightarrow \infty$ ,

so,  $\exists n_1$  s.t.  $\mathbb{E}[X] < \frac{n}{4}$  for all  $n \geq n_1$ .

Then, by Markov Ineq.,  $\mathbb{P}[X \geq \frac{n}{2}] < \frac{1}{2} < 1$ .

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $girth(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$ .

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$ .

For  $n \geq n_1$ ,  $P[X \geq \frac{n}{2}] < 1$  where  $X = \#$  cycles of length  $< g$ .

Next, we bound  $\alpha(G(n, p))$ . Set  $\beta = \lceil \frac{3}{p} \log n \rceil$ .

$$P[\alpha(G(n, p)) \geq \beta] \leq \binom{n}{\beta} (1-p)^{\binom{\beta}{2}} < ?$$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $\text{girth}(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

For  $n \geq n_1$ ,  $P[X \geq \frac{n}{2}] < 1$  where  $X = \#$  cycles of length  $< g$

Next, we bound  $\alpha(G(n, p))$ . Set  $\beta = \lceil \frac{3}{p} \log n \rceil$

$$P[\alpha(G(n, p)) \geq \beta] \leq \binom{n}{\beta} (1-p)^{\binom{\beta}{2}} < n^\beta e^{-p \binom{\beta}{2}} = e^{(\log n - p \frac{\beta-1}{2}) \beta}$$
$$\approx e^{(\log n - \frac{3}{2} \log n) \beta}$$

$\rightarrow 0$  as  $n \rightarrow \infty$

$\therefore \exists n_2$  s.t.  $P[\alpha(G(n, p)) \geq \beta] < \frac{1}{2}$  for  $n \geq n_2$ .

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $\text{girth}(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

For  $n \geq n_1$ ,  $P[X \geq \frac{n}{2}] < 1$  where  $X = \#$  cycles of length  $< g$

Next, we bound  $\alpha(G(n, p))$ . Set  $\beta = \lceil \frac{3}{p} \log n \rceil$

$$P[\alpha(G(n, p)) \geq \beta] \leq \binom{n}{\beta} (1-p)^{\binom{\beta}{2}} < n^\beta e^{-p \binom{\beta}{2}} = e^{(\log n - p \frac{\beta-1}{2}) \beta} \\ \approx e^{(\log n - \frac{3}{2} \log n) \beta} \\ \rightarrow 0 \text{ as } n \rightarrow \infty$$

$\therefore \exists n_2$  s.t.  $P[\alpha(G(n, p)) \geq \beta] < \frac{1}{2}$  for  $n \geq n_2$ .

If  $n \geq \max\{n_1, n_2\}$  then  $\exists G$  on  $n$ -vertices with  $X < \frac{n}{2}$  &  $\alpha(G) < \beta$



Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $\text{girth}(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

For  $n \geq n_1$ ,  $P[X \geq n/2] < 1$  where  $X = \#$  cycles of length  $< g$

Next, we bound  $\alpha(G(n, p))$ . Set  $\beta = \lceil \frac{3}{p} \log n \rceil$

$$P[\alpha(G(n, p)) \geq \beta] \leq \binom{n}{\beta} (1-p)^{\binom{\beta}{2}} < n^\beta e^{-p \binom{\beta}{2}} = e^{(\log n - p \frac{\beta(\beta-1)}{2}) \beta}$$
$$\approx e^{(\log n - \frac{3}{2} \log n) \beta}$$

$\rightarrow 0$  as  $n \rightarrow \infty$

$\therefore \exists n_2$  s.t.  $P[\alpha(G(n, p)) \geq \beta] < \frac{1}{2}$  for  $n \geq n_2$ .

If  $n \geq \max\{n_1, n_2\}$  then  $\exists G$  on  $n$ -vertices with  $X < n/2$  &  $\alpha(G) < \beta$

Remove a vertex from each of the  $X$  shortest cycles.

We get a graph  $G'$  on at least  $n/2$  vertices with  $\text{girth}(G') > g$  &  $\alpha(G') < \beta$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  
girth  $(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

For  $n \geq n_1$ ,  $P[X \geq \frac{n}{2}] < \frac{1}{2}$  where  $X = \#$  cycles of length  $< g$

For  $n \geq n_2$ ,  $P[\alpha(G(n, p)) \geq \beta] < \frac{1}{2}$  where  $\beta = \lceil \frac{3}{p} \log n \rceil$

$\exists$  graph  $G'$  on at least  $\frac{n}{2}$  vertices with girth  $\geq g$  &  $\alpha(G') < \beta$ .

$$\Rightarrow \chi(G') \geq \frac{n/2}{\beta} > \frac{pn}{6 \log n} = \frac{n^t}{6 \log n} \rightarrow \infty \text{ as } n \rightarrow \infty.$$

$\therefore \exists n_3$  s.t.  $\chi(G') > k$  for  $n \geq n_3$

Thm [Erdős 1959] Given  $k \geq 3, g \geq 3$ .  $\exists$  graph  $G$  with  $\text{girth}(G) \geq g$  and  $\chi(G) \geq k$ .

Proof Generate a random graph  $G(n, p)$

Let  $p = n^{t-1}$  where  $t < 1/g$ , say  $t = 1/2g$

For  $n \geq n_1$ ,  $P[X \geq \frac{n}{2}] < \frac{1}{2}$  where  $X = \# \text{cycles of length} < g$

For  $n \geq n_2$ ,  $P[\alpha(G(n, p)) \geq \beta] < \frac{1}{2}$  where  $\beta = \lceil \frac{3}{p} \log n \rceil$

$\exists$  graph  $G'$  on at least  $\frac{n}{2}$  vertices with  $\text{girth} \geq g$  &  $\alpha(G') < \beta$ .

$$\Rightarrow \chi(G') \geq \frac{n/2}{\beta} > \frac{pn}{6 \log n} = \frac{n^t}{6 \log n} \rightarrow \infty \text{ as } n \rightarrow \infty.$$

$\therefore \exists n_3$  s.t.  $\chi(G') > k$  for  $n \geq n_3$

So, if we choose  $n \geq \max\{n_1, n_2, n_3\}$

then  $\exists$  graph  $G$  with  $\text{girth} \geq g$  &  $\chi(G) \geq k$   
(on  $\geq \frac{n}{2}$  vertices) ■

Typical Probabilistic argument consists of defining certain "bad events"  $E_1, \dots, E_k$  that we want to avoid & then showing that the probability of doing so is positive.

→ If all  $P[E_i]$  are small, say  $\sum_i P[E_i] < 1$  then O.K.

Typical Probabilistic argument consists of defining certain "bad events"  $E_1, \dots, E_k$  that we want to avoid & then showing that the probability of doing so is positive.

→ If all  $P[E_i]$  are small, say  $\sum_i P[E_i] < 1$  then O.K.

→ If all  $E_i$  are independent then  $P[\text{none of } E_i \text{ occur}]$  is  $\prod_{i=1}^k (1 - P[E_i])$  which is  $> 0$  if all  $P[E_i] < 1$

But what if  $E_i$  are dependent?

Typical Probabilistic argument consists of defining certain "bad events"  $E_1, \dots, E_k$  that we want to avoid & then showing that the probability of doing so is positive.

→ If all  $P[E_i]$  are small, say  $\sum_i P[E_i] < 1$  then O.K.

→ If all  $E_i$  are independent then  $P[\text{none of } E_i \text{ occur}]$  is  $\prod_{i=1}^k (1 - P[E_i])$  which is  $> 0$  if all  $P[E_i] < 1$

But what if  $E_i$  are dependent?

Lovasz Local Lemma Let  $E_1, \dots, E_k$  be events with  $P[E_i] \leq p \forall i$ .

Suppose each  $E_i$  is independent of all other  $E_j$  except for at most  $d$  of them.

Then,  $e p (d+1) < 1 \Rightarrow P[\text{none of } E_i \text{ occur}] > 0$

Lovasz Local Lemma Let  $E_1, \dots, E_k$  be events with  $P[E_i] \leq p \forall i$ . Suppose each  $E_i$  is independent of all other  $E_j$  except for at most  $d$  of them. Then,  $e p(d+1) < 1 \Rightarrow P[\bigwedge_{i=1}^k \bar{E}_i] > 0$ .

Lets apply this to  $R(k, k)$ .

Theorem (Spencer 1977) If  $e \left( \binom{k}{2} \binom{n}{k-2} + 1 \right) 2^{1 - \binom{k}{2}} < 1$  then  $R(k, k) > n$ .

Proof

Lovasz Local Lemma Let  $E_1, \dots, E_k$  be events with  $P[E_i] \leq p \forall i$ . Suppose each  $E_i$  is independent of all other  $E_j$  except for at most  $d$  of them. Then,  $ep(d+1) < 1 \Rightarrow P[\bigwedge_{i=1}^k \bar{E}_i] > 0$ .

Lets apply this to  $R(k, k)$ .

Theorem (Spencer 1977) If  $e\left(\binom{k}{2}\binom{n}{k-2} + 1\right)2^{1-\binom{k}{2}} < 1$  then  $R(k, k) > n$ .

Proof Random 2-color edges of  $K_n$  (independently Red/Blue w. prob.  $\frac{1}{2}$ )  
For each  $S \in \binom{V(K_n)}{k}$ , let  $A_S$  be the event  $S$  induces a monochromatic  $K_k$ .

$$P[A_S] = 2^{1-\binom{k}{2}} = p.$$

$A_S$  is independent of all  $A_{S'}$  unless



Lovasz Local Lemma Let  $E_1, \dots, E_k$  be events with  $P[E_i] \leq p \forall i$ . Suppose each  $E_i$  is independent of all other  $E_j$  except for at most  $d$  of them. Then,  $ep(d+1) < 1 \Rightarrow P[\bigwedge_{i=1}^k \bar{E}_i] > 0$ .

Lets apply this to  $R(k, k)$ .

Theorem (Spencer 1977) If  $e\left(\binom{k}{2} \binom{n}{k-2} + 1\right) 2^{1-\binom{k}{2}} < 1$  then  $R(k, k) > n$ .

Proof Random 2-color edges of  $K_n$  (independently Red/Blue w. prob.  $\frac{1}{2}$ )  
For each  $S \in \binom{V(K_n)}{k}$ , let  $A_S$  be the event  $S$  induces a monochromatic  $K_k$ .

$$P[A_S] = 2^{1-\binom{k}{2}} = p.$$

$A_S$  is independent of all  $A_{S'}$  unless  $|S \cap S'| \geq 2$ .

For each  $S$ , there are at most  $\binom{k}{2} \binom{n}{k-2} = d$  such choices of  $S'$ .

Apply LLL to these events with  $p$  &  $d$  as above.

$\therefore$  with positive probability we have a 2-coloring of  $E(K_n)$  with no monochromatic  $K_k$   $\blacksquare$

## Our lower bounds of $R(k, k)$

• Union Bound  $(\frac{1}{e\sqrt{2}} + o(1)) k 2^{k/2}$

• Alteration  $(\frac{1}{e} + o(1)) k 2^{k/2}$

• LLL  $(\frac{\sqrt{2}}{e} + o(1)) k 2^{k/2}$

By Optimizing the choice of  $n$  in  
" $e \binom{k}{2} \binom{n}{k-2} + 1 \stackrel{d}{\leq} \underbrace{2^{1-\binom{k}{2}}}_p < 1 \Rightarrow R(k, k) > n$ "

Fix  $k$ , so  $p$  is fixed.

Make  $n$  small, so  $d$  is small enough.

$$d < \frac{k^2}{2} \left(\frac{ne}{k-2}\right)^{k-2} < \frac{1}{ep} = \frac{1}{2e} 2^{k/2} (2^{k/2})^{k-2}$$

want  $n \leq c \frac{\sqrt{2}}{e} k 2^{k/2}$

where  $c = \left(\frac{2}{ek^2}\right)^{1/(k-2)} \frac{k-2}{k}$

⋮

## Setup for Lovász Local Lemma:

Defn An event  $B$  is mutually independent of events  $A_1, \dots, A_k$

$$\text{if } P[B | C_{i_1} \wedge C_{i_2} \wedge \dots \wedge C_{i_\ell}] = P[B] \quad \forall \ell \quad \forall 1 \leq i_1 < i_2 < \dots < i_\ell \leq k.$$

where each  $C_{i_j}$  is either  $A_{i_j}$  or  $\overline{A_{i_j}}$ .

## Setup for Lovász Local Lemma:

Defn An event  $B$  is mutually independent of events  $A_1, \dots, A_k$

if  $P[B | C_{i_1} \wedge C_{i_2} \wedge \dots \wedge C_{i_\ell}] = P[B] \quad \forall \ell \quad \forall 1 \leq i_1 < i_2 < \dots < i_\ell \leq k.$

where each  $C_{i_j}$  is either  $A_{i_j}$  or  $\overline{A_{i_j}}$ .

Defn  $G$  is said to be a dependency graph of events  $A_1, \dots, A_n$

if  $V(G) = [n]$  and for each  $i$ ,  $A_i$  is mutually independent of all events  $A_j$  s.t.  $ij \notin E(G)$ .

## Setup for Lovász Local Lemma:

Defn An event  $B$  is mutually independent of events  $A_1, \dots, A_k$

$$\& \quad \mathbb{P}[B | C_{i_1} \wedge C_{i_2} \wedge \dots \wedge C_{i_\ell}] = \mathbb{P}[B] \quad \forall \ell \quad \forall 1 \leq i_1 < i_2 < \dots < i_\ell \leq k.$$

where each  $C_{i_j}$  is either  $A_{i_j}$  or  $\overline{A_{i_j}}$ .

Defn  $G$  is said to be a dependency graph of events  $A_1, \dots, A_n$

if  $V(G) = [n]$  and for each  $i$ ,  $A_i$  is mutually independent of all events  $A_j$  s.t.  $ij \notin E(G)$ .

Caution! mutual independence  $\neq$  pairwise independence.

e.g. Pick  $x_1, x_2, x_3 \in \{0, 1\}$  uniformly & independently at random.

For  $i=1, 2, 3$ , let  $A_i$  be event  $x_{i+1} + x_{i+2} = 0 \pmod{2}$  (indices  $\pmod{3}$ )

Then these events are pairwise independent but not mutually ind.

$\overline{K_3}$  is not a valid dep. graph, but is a valid dep. graph.



However, in many (most?) applications of LLL, the underlying probability space is a product probability space i.e., it is based on a collection of independent random experiments and each event  $A_i$  is determined by a subcollection  $S_i$  of these independent experiments.

However, in many (most?) applications of LLL, the underlying probability space is a product probability space i.e., it is based on a collection of independent random experiments and each event  $A_i$  is determined by a subcollection  $S_i$  of these independent experiments.

Mutual Independence Principle Let  $Z_1, Z_2, \dots, Z_m$  be independent experiments and  $A_1, \dots, A_n$  be events s.t. the occurrence of each  $A_i$  is determined by a subset  $S_i$  of  $Z_1, \dots, Z_m$ .  
If  $S_i$  is disjoint from  $S_{j_1}, \dots, S_{j_k}$  then  $A_i$  is mutually ind. of  $\{A_{j_1}, \dots, A_{j_k}\}$

However, in many (most?) applications of LLL, the underlying probability space is a product probability space i.e., it is based on a collection of independent random experiments and each event  $A_i$  is determined by a subcollection  $S_i$  of these independent experiments.

Mutual Independence Principle Let  $Z_1, Z_2, \dots, Z_m$  be independent experiments and  $A_1, \dots, A_n$  be events s.t. the occurrence of each  $A_i$  is determined by a subset  $S_i$  of  $Z_1, \dots, Z_m$ . If  $S_i$  is disjoint from  $S_{j_1}, \dots, S_{j_k}$  then  $A_i$  is mutually ind. of  $\{A_{j_1}, \dots, A_{j_k}\}$

In this set-up, a valid dependency graph can be formed by placing an edge  $i \leftrightarrow j$  based on  $|S_i \cap S_j|$ .

e.g. in our application of LLL to  $R(K, K)$ ,  $Z_1, \dots, Z_m$  are the independent "coin flips" used to color  $m$  edges of  $K_n$ .  $A_S$  is independent of all  $A_{S'}$  where  $|S \cap S'| \leq 1$ .



## Lovász Local Lemma, General Form

neighborhood in  
a dependency graph

Let  $A_1, \dots, A_n$  be events. For each  $i \in [n]$ , let  $N(i) \subseteq [n]$  be such that  $A_i$  is independent of  $\{A_j : j \notin \{i\} \cup N(i)\}$ .

if  $x_1, \dots, x_n \in [0, 1)$  satisfy

$$P[A_i] \leq x_i \prod_{j \in N(i)} (1 - x_j) \quad \forall i \in [n]$$

then with probability  $\geq \prod_{i=1}^n (1 - x_i)$ , none of the events  $A_i$  occur.

## Lovász Local Lemma, General Form

Let  $A_1, \dots, A_n$  be events. For each  $i \in [n]$ , let  $N(i) \subseteq [n]$  be such that  $A_i$  is independent of  $\{A_j : j \notin \{i\} \cup N(i)\}$ .

if  $x_1, \dots, x_n \in [0, 1)$  satisfy

$$P[A_i] \leq x_i \prod_{j \in N(i)} (1 - x_j) \quad \forall i \in [n]$$

then with probability  $\geq \prod_{i=1}^n (1 - x_i)$ , none of the events  $A_i$  occur.

## LLL, Symmetric Form

Let  $G$  be a dependency graph of events  $A_1, \dots, A_n$  with  $\Delta(u) \leq d$  and  $P[A_i] \leq p \quad \forall i$ . Then,

$$e p (d+1) < 1 \Rightarrow P\left[\bigwedge_{i=1}^n \bar{A}_i\right] > 0$$

Proof Set  $x_i = 1/(d+1) < 1 \quad \forall i$

Then  $x_i \prod_{j \in N(i)} (1 - x_j) \geq \frac{1}{d+1} \left(1 - \frac{1}{d+1}\right)^d > \frac{1}{(d+1)e} \geq p$

← neighborhood in dep. graph.

## Lovász Local Lemma, General Form

Let  $A_1, \dots, A_n$  be events. For each  $i \in [n]$ , let  $N(i) \subseteq [n]$  be such that  $A_i$  is independent of  $\{A_j : j \notin \{i\} \cup N(i)\}$ .

if  $x_1, \dots, x_n \in [0, 1)$  satisfy

$$P[A_i] \leq x_i \prod_{j \in N(i)} (1 - x_j) \quad \forall i \in [n]$$

then with probability  $\geq \prod_{i=1}^n (1 - x_i)$ , none of the events  $A_i$  occur.

## LLL, Symmetric Form

Let  $G$  be a dependency graph of events  $A_1, \dots, A_n$  with  $\Delta(u) \leq d$  and  $P[A_i] \leq p \quad \forall i$ . Then,

$$e p (d+1) < 1 \Rightarrow P\left[\bigwedge_{i=1}^n \bar{A}_i\right] > 0$$

$e$  is sharp; cannot be replaced by any smaller constant  
often, checking  $4p(d+1) < 1$  is enough.